Now. 17th.
\nHypothesis testing.
\n1. Point asfinaction
\n2. Connection with hypothesis testing.
\n3. Section 10.8, 10.9 (Small sample)
\nSection 10.8, 10.9 (Small sample)
\nSeeting: Population distribution: X ~ Tx. unknown
\nSamples: X,..., X. 236 Fx.
$$
\mu=0.87
$$
, $\sigma=Var X$
\nEstimation. $\bar{\mu} = \bar{\chi} = \frac{\Sigma X_i}{n}$, $\bar{\sigma}^2 = -\frac{1}{1} \Sigma (X_i \cdot \bar{X})^2$.
\nTry to figure out the distribution of the estimator, which is the core of estimation theory.
\n10. Fx is normal: M μ, σ^2).
\nNormal distribution has size properties.
\n $\bar{X} \sim M \mu, \frac{\sigma^2}{n}$, $\frac{\overline{E} \bar{\mu}}{\overline{V}} = \frac{N \sigma \bar{\mu}}{1}$
\n $\frac{M \cdot \overline{D}^2}{\sigma^2} \sim \chi^2_{n-1}$, $\frac{\overline{E} \bar{\mu}}{\overline{V}} = \frac{N \cdot \mu \cdot \overline{D}}{N} = \frac{N \cdot \mu \cdot \overline{D}^2}{\sqrt{1 - \frac{\overline{D}}{N}}}$
\n $\frac{\overline{X} \cdot \mu}{\overline{Y} \sqrt{n}} = \frac{(\overline{X} \cdot \mu) \overline{X} \cdot \mu}{\overline{Y} \sqrt{n}} \sim \frac{\overline{X} \cdot \mu}{\overline{Y} \sqrt{n}} = \frac{(\overline{X} \cdot \mu) \overline{X} \cdot \mu}{\overline{Y} \sqrt{n}}$

Sufficiency.

A Small-Sample Test for μ

Assumptions: Y_1, Y_2, \ldots, Y_n constitute a random sample from a normal distribution with $E(Y_i) = \mu$. $H_0: \mu = \mu_0.$

 $H_a: \begin{cases} \mu > \mu_0 & \text{(upper-tail alternative)}.\\ \mu < \mu_0 & \text{(lower-tail alternative)}.\\ \mu \neq \mu_0 & \text{(two-tailed alternative)}. \end{cases}$ L $\mu \neq \mu_0$

Test statistic: $T = \frac{\overline{Y} - \mu_0}{S/\sqrt{n}}$.

Rejection region: $\begin{cases} t > t_\alpha & \text{(upper-tail RR)} \\ t < -t_\alpha & \text{(lower-tail RR)} \\ |t| > t_{\alpha/2} & \text{(two-tailed RR)} \end{cases}$.

Small-Sample Tests for Comparing Two Population Means

Assumptions: Independent samples from normal distributions with $\sigma_1^2 = \sigma_2^2$. $H_0: \mu_1 - \mu_2 = D_0.$
 $H_a: \begin{cases} \mu_1 - \mu_2 > D_0 \\ \mu_1 - \mu_2 < D_0 \end{cases}$ (upper-tail alternative).
 $\mu_1 - \mu_2 \neq D_0$ (lower-tail alternative). Test statistic: $T = \frac{\overline{Y}_1 - \overline{Y}_2 - D_0}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$, where $S_p = \sqrt{\frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2}{n_1 + n_2 - 2}}$.

Rejection region: $\begin{cases} t > t_\alpha & \text{(upper-tail RR).} \\ t < -t_\alpha & \text{(lower-tail RR).} \\ |t| > t_{\alpha/2} & \text{(two-tailed RR).} \end{cases}$